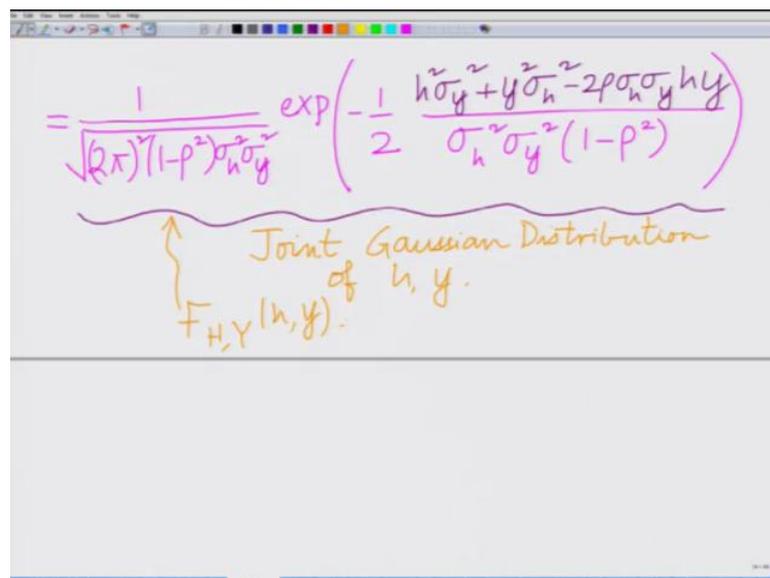


**Bayesian/MMSE Estimation for MIMO/OFDM Wireless Communications**  
**Prof. Aditya K. Jagannatham**  
**Department of Electrical Engineering**  
**Indian Institute of Technology, Kanpur**

**Lecture - 04**  
**Derivation of Minimum Mean Squared Error (MMSE) Estimate for Gaussian Parameter – part II**

Hello welcome to another module in this massive open online course of Bayesian MMSE estimation for wireless communication. So, currently we are looking at the m MSE estimate of the parameter h and, from an observation y when the parameter h and observation are jointly Gaussian, and towards this end we have derived the joint density of this h parameter h and the observation y, and that is given by this quantity.

(Refer Slide Time: 00:43)



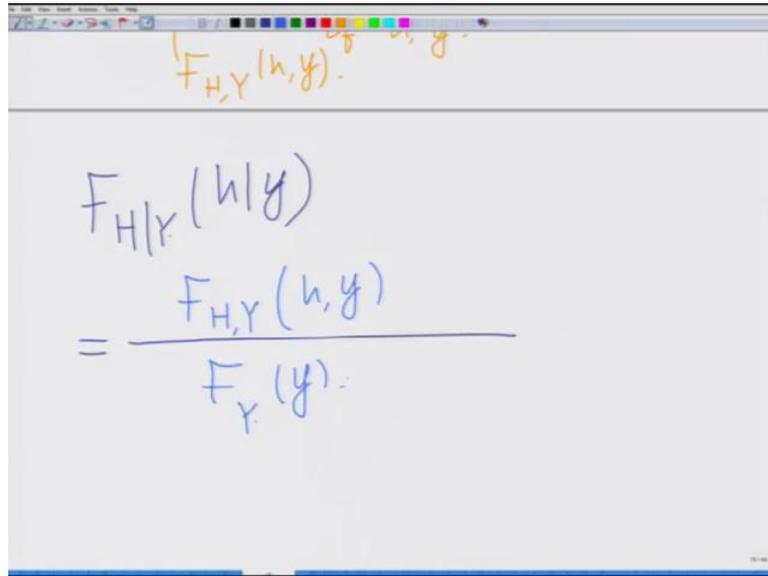
The image shows a handwritten mathematical formula on a whiteboard. The formula is:

$$= \frac{1}{\sqrt{(2\pi)^2 (1-\rho^2) \sigma_h^2 \sigma_y^2}} \exp\left(-\frac{1}{2} \frac{h^2 \sigma_y^2 + y^2 \sigma_h^2 - 2\rho \sigma_h \sigma_y h y}{\sigma_h^2 \sigma_y^2 (1-\rho^2)}\right)$$

Below the formula, there is a handwritten note in orange: "Joint Gaussian Distribution of h, y." with a bracket pointing to the formula. Below that, it says "F\_{H,Y}(h,y)." with an arrow pointing to the formula.

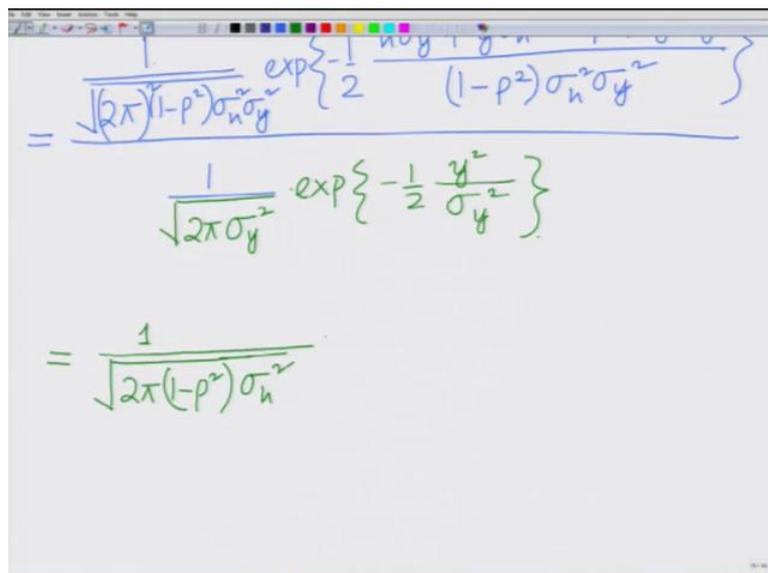
Given by this quantity that is shown here, that is done by previous module that is the joint probability density - The joint probability density function of the parameter y of the parameter h, and the observation y in the parameter h, and the observation y are the jointly Gaussian.

(Refer Slide Time: 01:06)


$$F_{H|Y}(h|y) = \frac{F_{H,Y}(h,y)}{F_Y(y)}$$

So, now as said before we need the conditional probability density function. I mean the conditional probability density function of in the parameter  $h$ , given the observation  $y$ , which is basically, as we have already said this is the joint probability density function of the parameter  $h$  observation  $y$  divided by the marginal probability density function of the observation  $y$ . And this now I am going to simplify by writing the exact expression for each.

(Refer Slide Time: 01:35)


$$= \frac{1}{\sqrt{2\pi(1-\rho^2)\sigma_h^2}} \exp\left\{-\frac{1}{2} \frac{(h - \rho y)^2}{(1-\rho^2)\sigma_h^2}\right\}$$
$$= \frac{1}{\sqrt{2\pi\sigma_h^2}} \exp\left\{-\frac{1}{2} \frac{y^2}{\sigma_y^2}\right\}$$
$$= \frac{1}{\sqrt{2\pi(1-\rho^2)\sigma_h^2}}$$

So, first is the joint probability density function that we have already derived the expressions. so that is square root of  $2\pi$  whole square  $1 - \rho^2$ , that is your determinant  $r$ , which is  $\sigma_h^2 \sigma_y^2 e^{-\frac{1}{2}(1 - \rho^2) \left( \frac{h^2 \sigma_y^2 + y^2 \sigma_h^2 - 2\rho \sigma_h \sigma_y h y}{\sigma_h^2 \sigma_y^2} \right)}$  divided by the marginal probability density function of  $y$  the observation  $y$ .

We said the marginal probability density function of the observation  $y$ ,  $y$  is Gaussian land of variable with means 0 variable  $\sigma_h^2$ ; therefore, this we all ready derive this marginal probability density function of the observation  $y$ , this is given as  $\frac{1}{\sqrt{2\pi} \sigma_y} e^{-\frac{y^2}{2\sigma_y^2}}$ . And therefore, now I can simplify this first letter, simplify this quantity that is the multiply factor, this is under root  $2\pi$   $1 - \rho^2$  whole square  $\sigma_h^2$  and  $\sigma_y^2$  divided by  $\sigma_h^2$ .

(Refer Slide Time: 03:56)

$$= \frac{1}{\sqrt{2\pi(1-\rho^2)\sigma_h^2}} \times \exp\left\{-\frac{1}{2}\left(\frac{h^2\sigma_y^2 + y^2\sigma_h^2 - 2\rho\sigma_h\sigma_y h y}{(1-\rho^2)\sigma_h^2\sigma_y^2} - \frac{y^2}{\sigma_y^2}\right)\right\}$$

quantity in the exponent.

So, this gives  $1 - \rho^2$  times, it gives  $1 - \rho^2$  times  $\sigma_h^2$  into multiplied by. This is a large term minus half of first, which is basically your  $1 - \rho^2$   $\sigma_h^2 \sigma_y^2$  plus  $y^2 \sigma_h^2$  minus  $2\rho \sigma_h \sigma_y h y$  minus this quantity, minus  $y^2$  divided by  $\sigma_y^2$  all right. So, this is the quantity in the exponent. So, this is multiplied by this is the multiplication

symbol, and this is the quantity in the exponent. So, this is the quantity in the exponent; e to the power of minus half times this quantity.

(Refer Slide Time: 05:20)

$$\frac{h^2 \sigma_y^2 + y^2 \sigma_h^2 - 2\rho h y \sigma_h \sigma_y}{(1-\rho^2)\sigma_h^2 \sigma_y^2} - \frac{\sigma}{\sigma_y^2}$$

$$= \frac{h^2 \sigma_y^2 + y^2 \sigma_h^2 - 2\rho h y \sigma_h \sigma_y - y^2 (1-\rho^2) \sigma_h^2}{(1-\rho^2)\sigma_h^2 \sigma_y^2}$$

$$= \frac{h^2 \sigma_y^2 + y^2 \rho^2 \sigma_h^2 - 2\rho h y \sigma_h \sigma_y}{(1-\rho^2)\sigma_h^2 \sigma_y^2}$$

Now, let us simplify this quantity in the exponents. So, I want to first start by simplifying this quantity in the exponent, and then we can simplify the rest of the expression later. And this quantity in the exponent that can be simplified as follows. Let me just write it down, it is a bit cumbersome, but this is fundamentally important in MMSE Estimation. So, it is worth going through this derivation, at least once; that is  $h^2 \sigma_y^2 + y^2 \sigma_h^2 - 2\rho h y \sigma_h \sigma_y$  divided by  $1 - \rho^2 \sigma_h^2 \sigma_y^2$  minus  $y^2 \sigma_h^2$  divided by  $\sigma_y^2$ .

And this is then I can combine these two  $1 - \rho^2 \sigma_h^2 \sigma_y^2$  square  $h^2 \sigma_y^2 + y^2 \sigma_h^2 - 2\rho h y \sigma_h \sigma_y$  plus  $y^2 \sigma_h^2$  minus  $y^2 \sigma_h^2$  into  $1 - \rho^2 \sigma_h^2 \sigma_y^2$ ; that is the term I will get over here. Now observe this quantity I have  $y^2 \sigma_h^2$ , and if I expand this I will have a minus  $y^2 \sigma_h^2$ . I will have a minus  $y^2 \sigma_h^2$  coming from these two. So, that those two terms. So, this term will cancel with this part, and what I am going to be left with interestingly if you observe this is  $h^2 \sigma_y^2$  square. Something very simple this  $y^2 \sigma_h^2$  cancels, what I am left with is  $y^2 \rho^2 \sigma_h^2$  square. So, I left with  $y^2 \rho^2 \sigma_h^2$  square minus  $2\rho h y \sigma_h \sigma_y$  plus  $1 - \rho^2 \sigma_h^2 \sigma_y^2$  square.

(Refer Slide Time: 08:06)

$$\begin{aligned} &= \frac{0}{(1-\rho^2)\sigma_h^2\sigma_y^2} \\ &= \frac{h^2\sigma_y^2 + y^2\rho^2\sigma_h^2 - 2\rho\sigma_h\sigma_y h y}{(1-\rho^2)\sigma_h^2\sigma_y^2} \\ &= \frac{(h\sigma_y - \rho\sigma_h y)^2}{(1-\rho^2)\sigma_h^2\sigma_y^2} = \frac{\left(h - \rho\frac{\sigma_h}{\sigma_y}y\right)^2}{(1-\rho^2)\sigma_h^2} \end{aligned}$$

*Quantity in the exponent.*

And now if you look at this quantity in the numerator that is nothing, but, observe that this quantity is nothing, but  $h\sigma_y - \rho\sigma_h y$  whole square, because look at this is  $h^2\sigma_y^2 + \rho^2\sigma_h^2 y^2 - 2\rho\sigma_h\sigma_y h y$ . So, this is nothing, but  $h\sigma_y - \rho\sigma_h y$  whole square.

And once we have that the rest is very simple, this is  $1 - \rho^2\sigma_h^2\sigma_y^2$ . and now take  $\sigma_y^2$  out common from the denominator and numerator that is  $\sigma_y^2$  common from the numerator and denominator, what I am going to have is  $\left(h - \rho\frac{\sigma_h}{\sigma_y}y\right)^2$  divided by  $1 - \rho^2\sigma_h^2$ , because  $\sigma_h^2$  cancels from the numerator and denominator,  $1 - \rho^2\sigma_h^2$ . This is basically your quantity in the exponent. Quantity in the exponent will be  $e$  to the power of minus half times this. This is your quantity the exponent, therefore, now the conditional, putting all together. Remember we set out to derive the conditional probability density function of parameter  $h$  giving the observation  $y$ .

(Refer Slide Time: 09:50)

The image shows a whiteboard with handwritten mathematical formulas. At the top right, there is a note: "arranging in exponent." The main formula is the conditional probability density function  $F_{H|Y}(h|y)$ , which is equal to  $\frac{1}{\sqrt{2\pi(1-\rho^2)\sigma_h^2}}$  multiplied by an exponential term  $\exp\left\{-\frac{1}{2} \frac{\left(h - \rho \frac{\sigma_h}{\sigma_y} y\right)^2}{(1-\rho^2)\sigma_h^2}\right\}$ . Below this, two new parameters are defined:  $\tilde{\sigma}^2 = (1-\rho^2)\sigma_h^2$  and  $\tilde{\mu} = \rho \frac{\sigma_h}{\sigma_y} y$ .

Therefore now the conditional probability density function of the parameter  $h$  given the observation  $y$ , conditional probability density function this will be given as follows. This will be given as, this is the multiplying factor is  $1$  over square root  $2\pi$   $1$  minus  $\rho$  square times  $\sigma_h$  square  $e$  to the power of minus half. Now you can look at this  $e$  to the power of minus half  $h$   $e$  to the power minus half  $h$  minus  $\rho$   $\sigma_h$  by  $\sigma_y$   $y$  whole square divided by  $1$  minus  $\rho$  square  $\sigma_h$  square. So, this is the joint probability density function. Now, what I am going to do is, basically I am going to set this quantity as  $\sigma$  tilde; that is I am going to define two new quantity  $\sigma$  tilde equal or  $\sigma$  tilde square equals  $1$  minus  $\rho$  square  $\sigma_h$  square and  $\mu$  tilde equals this quantity over here. I am going to set  $\mu$  tilde,  $\mu$  tilde equals  $\rho$   $\sigma_h$  by  $\sigma_y$   $y$  times  $y$ .

(Refer Slide Time: 11:36)

The image shows a whiteboard with handwritten mathematical notes. At the top, the conditional probability density function is written as  $f_{H|Y}(h|y) = \frac{1}{\sqrt{2\pi\tilde{\sigma}^2}} e^{-\frac{1}{2} \frac{(h-\tilde{\mu})^2}{\tilde{\sigma}^2}}$ . A bracket under the left side of this equation points down to the notation  $H|Y \sim \mathcal{N}(\tilde{\mu}, \tilde{\sigma}^2)$ . To the right of the main equation, the text "Gaussian Distribution" is written, followed by "Mean =  $\tilde{\mu}$ " and "Variance =  $\tilde{\sigma}^2$ ".

And now you can see something very interesting. now if I write this joint, if I write this conditional probability density function what I am going to observe is that; f of h given y of h given y equals 1 over 1 over, what you are going to observe is something very simple, 1 over square root 2 pi sigma tilde square e raise to minus half e raise to e is nothing, but expand, e raise to minus half h minus mu tilde whole square minus mu tilde whole square. And look at this, this is again your sigma tilde square divided by sigma tilde square. And now if you look at this, you will realize that this is a Gaussian distribution with mean mu tilde and variance sigma tilde square, and that is the interesting aspect of this.

So, this is basically, again now when you see this, that this is you can see that this is basically a Gaussian distribution with mean equals mu tilde variance equals sigma tilde. Therefore, what you can say is that the posterior. remember this also known as the posterior, the conditional probability density function of the parameter h given the observation y this is also the a posterior probability density function. This is Gaussian, so h given y h given y is basically Gaussian distributed with mean mu tilde variance sigma tilde square.

(Refer Slide Time: 13:31)

The image shows a whiteboard with handwritten mathematical derivations. At the top, it states  $H|Y \sim \mathcal{N}(\tilde{\mu}, \tilde{\sigma}^2)$ . Below this, it shows  $E\{h|y\} = \tilde{\mu} = \rho \frac{\sigma_h}{\sigma_y} y$ . A red arrow labeled "MMSE Estimation" points from the expression  $E\{h|y\}$  to the boxed equation  $\hat{h} = E\{h|y\} = \rho \frac{\sigma_h}{\sigma_y} y$ . A green arrow points from the boxed equation to the text "MMSE Estimate of h given y for h, y jointly Gaussian."

And therefore, now you can clearly see expected value of  $h$  given  $y$ , and this is very straight forward, now you can see the expected value of  $h$  given  $y$  equals  $\tilde{\mu}$  which is nothing, but now you can see what is  $\tilde{\mu}$ .  $\tilde{\mu}$  is basically this quantity over here; that is your  $\rho \sigma_h$  by  $\sigma_y$  times  $y$ , so this is your  $\rho \sigma_h$  over  $\sigma_y$  times  $y$ . And indeed now we are found the MMSE estimate, because what is the MMSE estimate. MMSE estimate is the minimum mean squared error estimate, remember the optimal minimum mean squared error estimate; that is the estimate which minimizes the mean squared error, is nothing, but the expected value of the parameter  $h$  given the observation  $y$ . Therefore, this quantity  $\tilde{\mu}$  is nothing, but the MMSE estimate of  $h$  given  $y$  for that scenario when  $h$  and  $y$  are jointly Gaussian; that is something in mind.

So, this is basically your MMSE estimate. So, let us also write that down clearly. It requires some manipulation, but the end result is very simple and clear. So, this says that the MMSE estimate  $\hat{h}$  equals expected value of  $h$  given  $y$  equals  $\rho \sigma_h$  divided by  $\sigma_y$  times  $y$ , this is the MMSE estimate of. So, this is the MMSE estimate  $h$  given  $y$  for  $h$  comma  $y$ ; that is the parameter and the observation, when the parameter and observation are basically jointly Gaussian random variation all right. And now what is this point is  $\tilde{\sigma}^2$ , this quantity, this is your mean squared error; this is the mean squared error.

(Refer Slide Time: 16:08)

$$E\{ (h - \hat{h})^2 | y \}$$
$$= \frac{(1 - \rho^2) \sigma_h^2}{}$$

Mean Squared Error

Interesting Observation:

If  $\rho = 1, -1 \Rightarrow 1 - \rho^2 = 0$ .

$h = y$   $h = -y$

So, this is the mean squared error, what is this. This is the expected value of the parameter  $h$  minus  $\hat{h}$  whole square given of course, the observation  $y$ . And what is this quantity. This quantity you can see  $1$  minus  $\rho$  square times  $\sigma_h$  square. So, this is the mean squared error. And something very interesting can be observed from this mean squared error. What is the interest thing that we can observe from this mean squared error? Consider two particular cases, if. So, let us write this as interesting observation. Consider the simple case, if  $\rho$  equals  $1$ . Remember the correlation coefficient  $\rho$  we said can lie between minus  $1$  or  $1$ . Now correlation coefficient of  $1$  basically means  $y$  and  $h$ ; that is the observation  $y$  and the parameter  $h$  are statistically identical, because the correlation is perfect; that is a correlation is the maximum value of one.

On the other hand if  $\rho$  is minus  $1$  there also statistically identical in different way; that is  $h$  is basically minus  $y$  and  $y$  is minus  $h$ . So, basically in both these cases we see; that is correlation coefficient  $\rho$  is equal to  $1$ , or the correlation coefficient  $\rho$  is minus  $1$   $\rho$  square is equal to  $1$ . Therefore,  $1$  minus  $\rho$  square is  $0$  which means MSE is  $0$ , and that is naturally true because if the correlation coefficient is  $1$ , on observing  $y$  you know that  $h$  is identically equal to  $y$ . And when correlation coefficient is minus  $1$  on observing  $y$  you know that  $h$  is equal to minus  $y$ , so there is no error, in that sense the mean squared error is  $0$ . When there is perfect correlation between the observation  $y$  and the parameter  $h$ , the mean squared error is  $0$ , and that is the interesting implication from this derivation.

So, rho is equal to 1 or for that matter minus 1, we have this implies 1 minus rho square equal to 0. So, this case, what we can see is in this case h is basically equal to y or h is basically equal to minus y, h can be perfectly estimated.

(Refer Slide Time: 18:54)

$$\text{If } \rho = 0 \Rightarrow 1 - \rho^2 = 1$$

$$\Rightarrow \text{MSE} = \sigma_h^2$$

$$= \text{Initial variance of Parameter } h.$$

$$E\{h^2\} = \sigma_h^2 = r_{hh}$$

$$E\{y^2\} = \sigma_y^2 = r_{yy}$$

$$E\{hy\} = \rho \sigma_h \sigma_y = r_{hy}:$$

So, look we are saying is h can be, h can be perfectly estimated on y, since the correlation is minus 1 or 1. The correlation is 1 h is equal to y, if the correlation is minus 1 h is equal to y h is equal to minus 1. Now consider the other scenario, if rho is equal to zero; that is implies the correlation between h and y is 0 implies 1 minus rho square equals 1 implies MMSE equals what is the MSE, you can see 1 minus rho square is equal to 1 MSE rho is equal zero, MSE simply sigma h square the mean squared error simply sigma; that is initial variance which is equal to, which is equal to initial variance of the parameter h. it means having the, because the correlation between the observation y and h is zero, the observation y does not convey any information about a.

So, observation y cannot be used in any meaningful sense, to get an estimate on the unknown parameter h. Therefore, the variance of h; that is the there is the parameter h is sigma h square to begin with, even after making the observation y the variance of h means that sigma h square. Therefore, having observed y does not make any impact, or does not convey any information about the unknown parameter h. These are two extreme cases which interesting insights into the behavior of this MSE estimate. And now we can see for different values of rho between 0 and 1, and the variance 0 can go all the way

from  $\sigma_h^2$ ; that is when  $\rho$  is equal to 0 the variance is  $\sigma_h^2$ . When  $\rho$  is equal to 1 the variance is zero, and between when  $\rho$  increases, as  $\rho$  increases from 0 to 1 the variance decreases, or the MSE decreases from  $\sigma_h^2$  to zero. So, that is the interesting observation that one can make from the MSE; the mean squared error per say, of this MMSE estimate that we derived from the jointly Gaussian scenario.

So, now let us also simplify this MMSE estimate and the variance to give alternative expression, which are both convenient. So, we have already said, what have we already said? We have already said if you go back to our nomenclature, and if you look at our definitions. From our definitions what we have, is basically, we have that expected  $h$  square equals  $\sigma_h^2$ . Let us denote this by  $r_{hh}$ , expected  $y$  square equals  $\sigma_y^2$ , let us denote this by  $r_{yy}$ . just denoting this by the variance by  $r_{hh}$ , and the variance of  $y$  by  $r_{yy}$ , and the cross correlation between  $h$  and  $y$ . This we can denote by expected  $h$  comma  $y$  by why we denote by  $r_{hy}$ . So, let me just repeat those definitions over here. So, that it becomes convenient. let us define your  $\sigma_h^2$  as expected value of the parameter  $\sigma_h^2$  equals expected  $h$  square  $\sigma_h^2$  equals  $r_{hh}$ . Let us define expected value of  $y$  square equals  $\sigma_y^2$  as  $r_{yy}$ . And let us define the expected value of  $h$   $y$  which is equal to  $\rho \sigma_h \sigma_y$  as  $r_{hy}$ .

(Refer Slide Time: 23:15)

$$\hat{h} = \frac{\rho \sigma_h}{\sigma_y} y \quad r_{hy}$$

$$= \frac{\rho \sigma_h \sigma_y}{\sigma_y^2} y \quad r_{yy}$$

$$\hat{h} = r_{hy} r_{yy}^{-1} y$$

Now if you look at the MMSE estimate, the MMSE estimate is  $\hat{h}$  which is the expected value of  $h$  given  $y$ , which is also you can see for this  $\rho \sigma_h$  divided by  $\sigma_y$  times  $y$ . So, this is basically your  $\rho \sigma_h$  divided by  $\sigma_y$  times  $y$ . And now I can simplify this as follows; that is  $\rho \sigma_h \sigma_y$  multiplying and dividing by  $\sigma_y$  I have  $\rho \sigma_h \sigma_y$  divided by  $\sigma_y^2$  times  $y$ , which is now if you look at this, this quantity here, this is basically your  $r_{hy}$ , and this quantity in the denominator basically your  $r_{yy}$ . So, I can write this as something that is going to be helpful in our later discussion  $r_{hy}$  times  $1$  over  $r_{yy}$ , which is basically the  $r_{yy}$  inverse times  $y$ .

And this is a very interesting expression, because the same expression can be carried over also for the matrix scenario; that is when  $h$  and  $y$  are vectors. So, I can write the estimate  $\hat{h}$  as  $r_{hy}$  times  $r_{yy}$  inverse times  $y$ . Now of course, when  $h$  and  $y$  are vectors the interesting thing is I can replace  $r_{yy}$  by the covariance matrix of  $y$   $r_{yy}$  by the cross covariance matrix between the parameter vector  $h$  and the observation. So, this is the very interesting way to write this, because is a very convenient expression to remember, and that can also be used in various other much more complicated scenarios, such as vector scenario, when  $y$  is a vector when  $h$  is also a parameter vector. Although the expression that we have derived which is  $\rho \sigma_h$  divided by  $\sigma_y$  times  $y$  is equally valid, this is much more convenient and much more standard form to remember this expression.

(Refer Slide Time: 25:19)

The image shows a handwritten derivation of the MMSE estimate formula. It starts with the scalar case:  $MSE = (1 - \rho^2) \sigma_h^2$ . This is then expanded to  $= \sigma_h^2 - \rho^2 \sigma_n^2$ . The next step is  $= r_{hh} - \rho \sigma_h \times \rho \sigma_n$ . This is further simplified to  $= r_{hh} - \frac{r_{hy}}{\sigma_n \sigma_y} \cdot \frac{1}{\sigma_y^2} \rho \sigma_n \sigma_y$ . Finally, the vector case is boxed:  $MSE = r_{hh} - r_{hy} \cdot r_{yy}^{-1} \cdot r_{yh}$ . A note indicates that MMSE stands for minimum mean squared error.

$$\begin{aligned}
 MSE &= (1 - \rho^2) \sigma_h^2 \\
 &= \sigma_h^2 - \rho^2 \sigma_n^2 \\
 &= r_{hh} - \rho \sigma_h \times \rho \sigma_n \\
 &= r_{hh} - \frac{r_{hy}}{\sigma_n \sigma_y} \cdot \frac{1}{\sigma_y^2} \rho \sigma_n \sigma_y \\
 \text{MMSE (minimum mean squared error)} \\
 \boxed{MSE} &= r_{hh} - r_{hy} \cdot r_{yy}^{-1} \cdot r_{yh}
 \end{aligned}$$

And the expression for the MSE can be simplified as, so this is the expression for the MMSE estimate. And the expression for the MSE this can be simplified as follows. So, this is the expression for MMSE estimate and the expression for the MSE this can be simplified as follows this equal to  $1 - \rho^2 \sigma_h^2$ , which is basically your  $\sigma_h^2 - \rho^2 \sigma_h^2$ . Now of course, this quantity here we already know what this quantity is. This quantity is  $r^T y$ , so this is, I am sorry this quantity here is basically your  $r^T h$ . So, this is basically  $r^T h$  minus. Now this quantity I can write as  $\rho \sigma_h$  times  $\rho \sigma_h$ . Now, this quantity I will again manipulate as by multiplying and dividing by  $\sigma_y^2 \rho \sigma_h \sigma_y$  into  $1 / \sigma_y^2$  into  $\rho \sigma_h \sigma_y$ .

Now you can this quantity is  $r^T h$ , this quantity is here  $r^T y$ , and this quantity here I can write it as  $r^T h$ , because for the scalar scenario  $r^T h$  equals  $r^T y$  which might not to be same as the vector scenario. In fact,  $r^T h$  is  $r^T y$  transpose for real vector. So, this is  $r^T h$  into  $r^T y$  inverse into  $r^T y$ , and this is your convenient expression for the MMSE m s. In fact, this is not just the MMSE, but this is the MMSE this. in this case this is the minimum mean squared error, it is not just the MMSE, but for this particular scenario this is the MMSE; that is your actual minimum mean squared error. So, we have found two things; one, we have found the MMSE estimate; that is  $\rho \sigma_h \sigma_y$  divided by  $\sigma_y^2$  into  $y$  which is basically  $r^T h$  times  $r^T y$  inverse into  $y$ . And also we have found the actual value of this minimum mean squared error; that is the MMSE which is  $r^T h$  that is  $\sigma_h^2 - r^T h$  into  $r^T y$  inverse into  $r^T h$ , this is your means squared error for this scenario.

So, that basically completes this derivation where we have started with jointly Gaussian random variables  $h$  and  $y$ , where  $h$  is the random parameter,  $y$  is the random observation, we considered into the jointly Gaussian in nature. And following this we basically derived the conditional probability density function of  $h$  given  $y$ , using the joint probability density function of  $h$  and  $y$ , and the marginal probability density function of  $y$ , and basically from that we inferred the conditional probability, the conditional mean of the unknown parameter  $h$  given the observation  $y$  which is in fact, the MMSE estimate. And we have also found the mean squared error or rather in this case also the minimum mean squared error.

So, we will stop this module here, and we will continue the other aspects in subsequent modules.

Thank you very much.