

Applied multivariate Statistical Modeling
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Lecture - 10
Multivariate Normal distribution

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Multivariate Normal Distribution

$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2} \quad -\infty < x < \infty$$

pdf

$\mu, \sigma \rightarrow$ population parameters.

$x \rightarrow X = \begin{bmatrix} x_1 \\ x_2 \\ \vdots \\ x_p \end{bmatrix} \quad \mu \rightarrow \mu = \begin{bmatrix} \mu_1 \\ \mu_2 \\ \vdots \\ \mu_p \end{bmatrix} \quad \sigma^2 \rightarrow \Sigma = \begin{bmatrix} \sigma_{11} & \sigma_{12} & \dots & \sigma_{1p} \\ \sigma_{21} & \sigma_{22} & \dots & \sigma_{2p} \\ \vdots & \vdots & \ddots & \vdots \\ \vdots & \vdots & \dots & \sigma_{pp} \end{bmatrix}$

$f(x) \sim N_p(\mu, \Sigma)$

Today, we discuss multivariate normal distribution multivariate normal distribution, last class we have see the invariate normal distribution. You see the formula if x is a random variable then $2\pi\sigma^2 e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}$ minus infinite less than x less than plus infinite, so the p d f. Multivariate normal p d f is characterized by mu and sigma mu and sigma that is a population parameter. We want the counter part of p d f multivariate domain when x that invariate x is converted to X which is your X 1, X 2, X p and univariate mu is no longer univariate. It will be a very mean vector mu 1, mu 2, mu p.

Similarly, invariate sigma square will no longer be univariate, it will be a multivariate covariance matrix p cross p. So, when we want something by multivariate normal distribution, we want something which is f x in terms of N variable number p and mu vector and covariance matrix, how do you, how do you go about it and how to do it that is the discussion, today.

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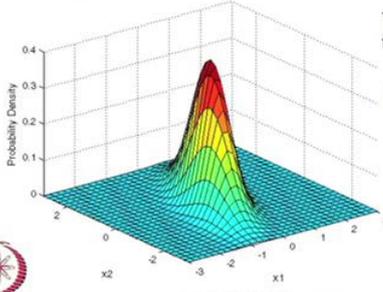
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Today, we will discuss bivariate normal density function multivariate normal density function and properties of multivariate normal density function. If time permits, we will go for statistical distance and constant density contours.

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An Example

A process is designed to produce laminar aluminium sheet of length X_1 and breadth X_2 . Let the deviation along X_1 and X_2 from target T_1 and T_2 are x_1 and x_2 , where X_1 , X_2 , x_1 and x_2 are random variables.



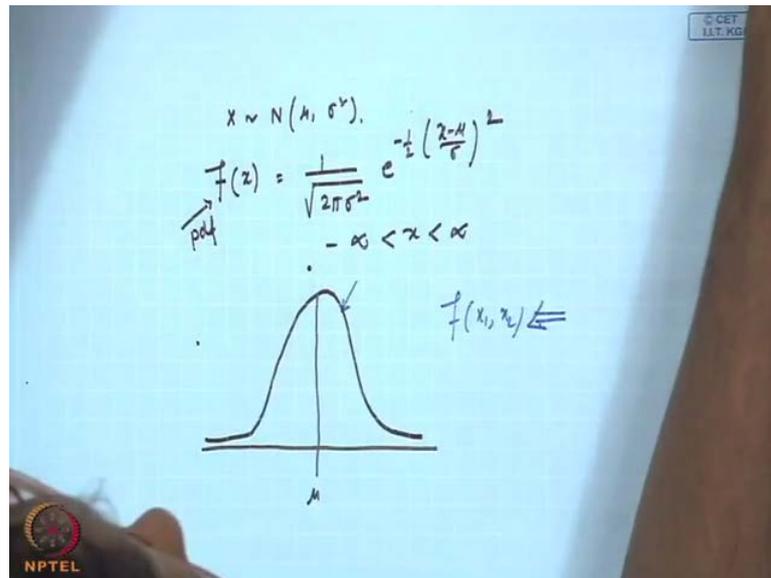
The pdf of x_1 and x_2 is shown in the figure left.

Here, $p = 2$ and pdf is $N_2(\mu, \Sigma)$.

$$X = \begin{bmatrix} x_1 \\ x_2 \end{bmatrix}$$
$$\mu = \begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix} \quad \Sigma = \begin{bmatrix} \sigma_1^2 & \sigma_{12} \\ \sigma_{12} & \sigma_2^2 \end{bmatrix}$$


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So, the univariate PDF is this you want to visualize its multivariate counterpart, so let us consider a bivariate case. You see this slide, in this slide you see that there are two variable X_1 and X_2 and probability density that is joint density that X_1 and X_2 X_1 and X_2 . So, this is what is given in figure, so you see that you are getting a bell shape, but in three dimensional you are getting because there first two dimension for the two variable values and third dimension are the density values.

If you take one more dimension, it is difficult you cannot visualize, suppose there are three variables with density, we cannot visualize pictorially. Now, as I told you our objective of the first part of today's lecture is we want to develop, this is our objective. So, in order to do so, we will follow a systematic, but simple path.

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$X_1, X_2 \leftarrow \text{bivariate Case}$
 $X = \begin{bmatrix} X_1 \\ X_2 \end{bmatrix}$
 $\mu = \begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix}$
 $\Sigma = \begin{bmatrix} \sigma_{11} & \sigma_{12} \\ \sigma_{12} & \sigma_{22} \end{bmatrix}$
 $\sigma_{12} = 0, \rho_{12} = 0$
 $f(x_1, x_2) = f(x_1) \cdot f(x_2)$
 $= \frac{1}{\sqrt{2\pi\sigma_1^2}} e^{-\frac{1}{2} \left(\frac{x_1 - \mu_1}{\sigma_1}\right)^2} \times \frac{1}{\sqrt{2\pi\sigma_2^2}} e^{-\frac{1}{2} \left(\frac{x_2 - \mu_2}{\sigma_2}\right)^2}$
 $= \frac{1}{(2\pi)^{j/2} (\sigma_1 \sigma_2)^{j/2}} e^{-\frac{1}{2} \left[\left(\frac{x_1 - \mu_1}{\sigma_1}\right)^2 + \left(\frac{x_2 - \mu_2}{\sigma_2}\right)^2 \right]}$
 $- \infty < x_j < \infty, j=1, 2$

Suppose, you think that you have two variables X_1 and X_2 which we are saying a bivariate case, so our X is X_1 and X_2 that is why my μ is that mini vector μ_1 and μ_2 . Your covariance matrix will be 2 by 2 σ_{11} , σ_{12} , σ_{12} , σ_{22} , I hope that there is no problem with you in this nomenclature. So, we assume something here, you see the slide here, in the slide you see the top figure that here you just scatter plot.

Now, what you can say about the two variables X_1 and X_2 seeing the scatter plot are they co related or there is no correlation, is it something like a circle you are getting or ellipse. There is no pattern you see that it is a ellipse type of thing, but there is no correlation. So, we want to simplify our derivation without correlation, so let me know what it means to say our σ_{12} is 0 or ρ_{12} is 0. If there is no correlation, then what will be the joint density?

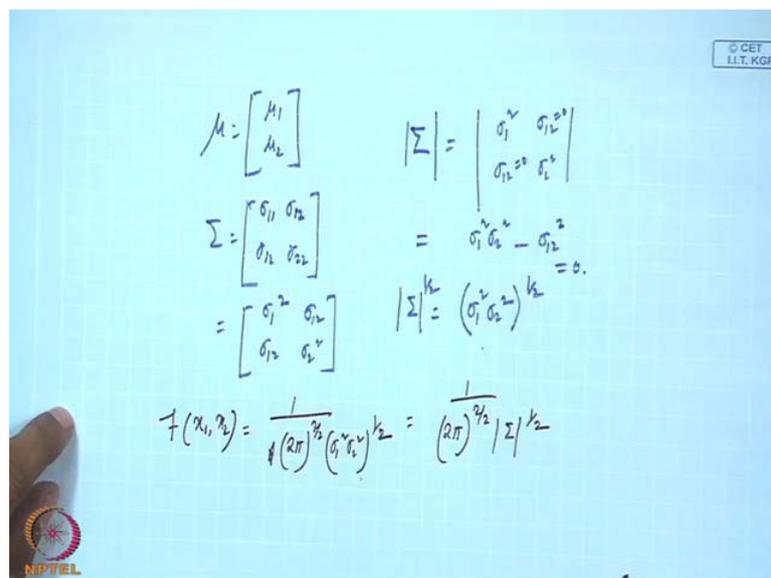
Suppose, X_1 and X_2 multiplication of the marginal density of the two, so X_1 cross X_2 , now all of us know that if X equal to 1 by root over 2 pi sigma 1 square as it is X_1 e to the power minus half X_1 minus μ_1 by sigma 1 square. Then, similarly for X_2 also, that is a second variable, you can write 2 pi sigma 2 square e to the power minus half X_2 minus μ_2 by sigma 2 square. You can write this and definitely here x_j less than greater than minus infinity to less than infinity to j equal to 1, 2 variable you have taken.

So, you can multiply these what you are getting, we are getting like this 1 by, that is one quantity equal to 2 pi, so 2 pi you are getting 2 by 2 pi into 2 pi square root 2 by 2. Then, another one what you are getting, sigma 1 square and sigma 2 square also, sigma 1 square sigma 2 square to the power half you are getting here. Then, I am coming to the exponent part e to the power minus half and all of us know that e to the power a into e 2 power b equal to e to the power a plus b.

So, we can write this one like this X_1 minus mu 1 by sigma 1 square plus X_2 minus mu 2 by sigma 2 squares, we can write this. So, essentially what is happening here that when I go for that univariate normal or bivariate normal with this with no dependence structure. You are having two component in the density function, one is the constant part another one is the exponent part; exponent means e to the power of something.

When I am making the joint distribution, here also you are having also two part this and this, the general structure for the multivariate normal distribution invariate that remain same what is the difference, difference will come in the two components and the values will be different. So, we found out that if X_1 and X_2 are independent, then our structure is like this, now let us see that we want to derive this constant part as well as exponent part from the population parameter.

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$$\mu = \begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix}$$

$$\Sigma = \begin{bmatrix} \sigma_{11} & \sigma_{12} \\ \sigma_{12} & \sigma_{22} \end{bmatrix}$$

$$= \begin{bmatrix} \sigma_1^2 & \sigma_{12} \\ \sigma_{12} & \sigma_2^2 \end{bmatrix}$$

$$|\Sigma| = \begin{vmatrix} \sigma_1^2 & \sigma_{12} \\ \sigma_{12} & \sigma_2^2 \end{vmatrix}$$

$$= \sigma_1^2 \sigma_2^2 - \sigma_{12}^2 = 0.$$

$$|\Sigma|^{1/2} = (\sigma_1^2 \sigma_2^2)^{1/2}$$

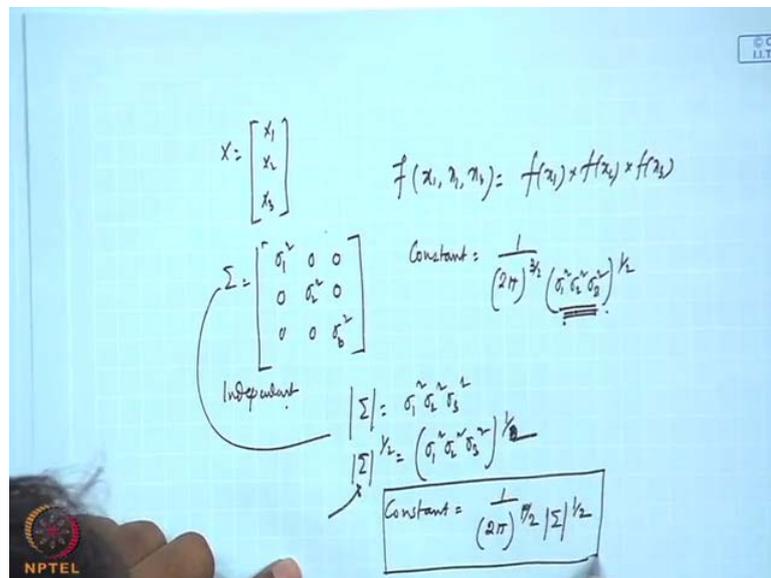
$$f(x_1, x_2) = \frac{1}{(2\pi)^2 (\sigma_1^2 \sigma_2^2)^{1/2}} = \frac{1}{(2\pi)^2 |\Sigma|^{1/2}}$$

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We have μ_1 into μ_2 μ equal to $\mu_1 \mu_2$ and we have σ is σ_{11} , σ_{12} , σ_{12} , σ_{22} , this one you can write. Now, σ_{11} equal to σ_{11}^2 σ_{12} σ_{12} σ_{22} square, so if I make something like this determinant of σ . Here, σ_{11}^2 , σ_{12} , σ_{12} , σ_{22}^2 , its determinant and you know that determinate will be this cross this minus this cross this. So, this one is $\sigma_{11}^2 \sigma_{22}^2$ minus σ_{12}^2 square, now you see that what we have assumed in the earlier demonstration.

We say σ_{12} equal to 0, just for the sake of simplicity we have taken that σ_{12} is 0, so if σ_{12} is 0, then determinant of σ is nothing but $\sigma_{11}^2 \sigma_{22}^2$. So, if I make square root of these, then this is the determinant of square root of the determinant and if this is the case. Then, the constant part what is the in case of our independent bivariate density function, we found out that constant part is $2\pi^2$ by 2 $\sigma_{11}^2 \sigma_{22}^2$ to the power half. Now, these I can write like these $2\pi^2$ by 2 determinant of covariance matrix to the power half.

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Now, suppose you have one more variable that means you have taken three variables, now like X is X_1 , X_2 and X_3 . You have to consider that your σ is σ_{11}^2 , 0 , 0 , 0 , σ_{22}^2 , 0 , 0 , 0 , σ_{33}^2 , we are assuming that all the variables are independent, then what will happen again $f(x_1, x_2, x_3)$ will be $f(x_1)$ into $f(x_2)$ into $f(x_3)$.

In the similar way you multiply, ultimately your constant term will be 1 by 2π , now three variables are there by 2 , then sigma 1 square, sigma 2 square, sigma 3 square to the power half. You see if you take determinant here, what are you getting here it will be sigma 1 square, sigma 2 square, sigma 3 square. So, that means determinant to the power half is sigma 1 square sigma 2 square sigma 3 square to the power of half.

So, if you now increase it to p variables, so ultimately your dimension will change and sigma to the power half will take care of one part of the constant. So, if I go by p variable, now my constant will become like this one by you see that, when there are two variables it is 2 by 2 when three variables 2π to the power of 3 by 2 .

So, when there are p variable, it will be p by 2 and whether it is two variable or three variable, three variable case. Ultimately, this quantity will be replaced by determinant of covariance matrix to power half, so with one assumption here that we are considering independent variable we proved this is the case, now what will happen to your constant exponent term.

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The image shows a handwritten derivation on a blue grid background. At the top right, there is a small logo for '© CET IIT KGP'. The main derivation starts with the expression:

$$\frac{1}{(2\pi)^{p/2} (\sigma_1^2 \sigma_2^2 \dots \sigma_p^2)^{1/2}} e^{-\frac{1}{2} \left[\left(\frac{x_1 - \mu_1}{\sigma_1} \right)^2 + \left(\frac{x_2 - \mu_2}{\sigma_2} \right)^2 \right]}$$

Below this, the term $(x - \mu)^T$ is written, and the function is defined as $f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2} \left(\frac{x - \mu}{\sigma} \right)^2}$. The exponent is then simplified as follows:

$$\text{Exponent} = -\frac{1}{2} \underbrace{(x - \mu)}_X \underbrace{(\sigma^2)^{-1}}_\Sigma \underbrace{(x - \mu)}_X = -\frac{1}{2} \left(\frac{x - \mu}{\sigma} \right)^2$$

An arrow points down to a boxed expression:

$$-\frac{1}{2} (x - \mu)^T \Sigma^{-1} (x - \mu)$$

At the bottom left, there is a small logo for 'NPTEL'.

So, in two variable cases we found out that the exponent term $X_1 - \mu_1$ by sigma 1 square plus $X_2 - \mu_2$ by sigma 2 squared. So, we say this is the exponent term this is the exponent term this portion is exponent term, so it is x you see that $X_1 - \mu_1$ on that is been subtracted divided by a standard deviation that is square.

Now, create one suppose x minus μ transpose, no I will explain from the invariate case that will be better, so invariate case $f(x)$ by $\frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(x-\mu)^2}$. So, your exponent is minus half, I am writing x minus μ sigma square to the power minus 1 x minus μ is minus half x minus μ by sigma square. You see x minus half is there minus half x minus μ x minus μ square divided by sigma square, so sigma square sigma to the power inverse.

Now, if you go for the multivariate case what will happen your x is replaced by X , μ is replaced by bold μ , sigma square will be replaced by sigma. Now, in matrix multiplication what will be the square transpose that matrix X transpose x that is the square term.

So, we are basically making here square, so we want this that is why what is meant to say in multivariate domain, the exponent can be written like this x minus μ transpose sigma square is replaced by sigma to the power minus 1 x minus μ . From univariate normal distribution, we have taken the exponent part and we are saying that if we go in same manner to the multivariate part our resultant quantity will be this for the exponent is it so?

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$$X = \begin{bmatrix} x_1 \\ x_2 \end{bmatrix} \quad \mu = \begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix} \quad \Sigma = \begin{bmatrix} \sigma_1^2 & 0 \\ 0 & \sigma_2^2 \end{bmatrix}$$

$$-\frac{1}{2} \begin{bmatrix} x_1 - \mu_1 & x_2 - \mu_2 \end{bmatrix} \begin{bmatrix} \sigma_1^2 & 0 \\ 0 & \sigma_2^2 \end{bmatrix}^{-1} \begin{bmatrix} x_1 - \mu_1 \\ x_2 - \mu_2 \end{bmatrix}$$

$$A = \begin{bmatrix} \sigma_1^2 & 0 \\ 0 & \sigma_2^2 \end{bmatrix} \quad A^{-1} = \begin{bmatrix} \frac{1}{\sigma_1^2} & 0 \\ 0 & \frac{1}{\sigma_2^2} \end{bmatrix} \quad |\Sigma| = \sigma_1^2 \sigma_2^2$$

$$A^{-1} = \frac{\text{Adj}(A)}{\det(A)}$$

$$= \frac{\text{Transpose of cofactor of } A}{\det A}$$

$$= a_{ij} = (-1)^{i+j} a_{ji}$$

$$\text{Cofactor of } \Sigma = \begin{bmatrix} (-1)^{11} \sigma_2^2 & 0 \\ 0 & \sigma_1^2 \end{bmatrix}$$

$$= \begin{bmatrix} \sigma_2^2 & 0 \\ 0 & \sigma_1^2 \end{bmatrix}$$

Now, can we do like the same thing for X equal to X_1 and X_2 variable case μ equal to μ_1 and μ_2 and sigma equal to you have taken already sigma 1 square, 0, 0, sigma 2 square because we are independent case. We want to prove first because we know under

this condition, what will be the distribution multivariate normal density function that is known.

So, then you write down minus half, so $x - \mu$ transpose, so that means this is $X_1 - \mu_1$ $X_2 - \mu_2$ because $x - \mu$ is $X_1 - \mu_1$ $X_2 - \mu_2$, 2 cross 1 it will be 1 cross 2. Now, your what do you want σ_1^2 , 0, 0, σ_2^2 square, this inverse then, so that is 2 cross 2 then $X_1 - \mu_1$, $X_2 - \mu_2$, this is your 2 cross 1.

So, what is the resultant quantity 1 cross 2, 2 cross 2, 2 cross 2, it is a 1 cross 1, this will give you 1 cross 2, this will give you 1 cross 1. We say density that exponential to the power this constant value, you will be getting some values density will be calculated. Now, what is the inverse, how to calculate the inverse suppose if A is a matrix like this a 1 1, a 1 2, a 2 1, a 2 2, how do you compute the inverse 1 by adjoint by determinant.

So, A inverse is adjoint of A by determinant of A, now adjoint is the transpose of the cofactors of A divided by determinant of A. So, this is the case our A is nothing but this one σ_1^2 , 0, 0, σ_2^2 square, which is what is our sigma. Now, determinant already we have seen the determinant is σ_1^2 and σ_2^2 square multiplied by these two. Now, what will be the cofactor of this cofactor is if you if you suppose I want to know cofactor of sigma.

Here, suppose in a case you see cofactor means suppose you want to see the cofactor of these then you have to cross this corresponding row and column what is left that is the cofactor, but the sign conversion will be there. So, that means cofactor means for a i j, the cofactor will be $(-1)^{i+j}$ and the remaining portion whatever the remaining portion remaining part of the matrix that will be the case. As you have take 2 by 2 cross, so ultimately what will happen one row and one column crossed means only one item will be left.

So, it is our case, then cofactor of these we can write first one is minus 1 to the power 1 plus 1 that is 1 plus 1 then what is remaining here, σ_2^2 squares. Suppose, you cross this and this σ_2^2 square will be there and see it is 0 and it will be also 0 and σ_1^2 square will be this. So, our cofactor is σ_2^2 square, 0, 0, σ_1^2 square what will be transpose same because these two element are 0.

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$$\begin{aligned} \text{Transpose of cofactor of } \Sigma &= \begin{bmatrix} \sigma_2^2 & 0 \\ 0 & \sigma_1^2 \end{bmatrix} \\ \Sigma^{-1} &= \frac{1}{\sigma_1^2 \sigma_2^2} \begin{bmatrix} \sigma_2^2 & 0 \\ 0 & \sigma_1^2 \end{bmatrix} \\ -\frac{1}{2} \begin{bmatrix} x_1 - \mu_1 & x_2 - \mu_2 \end{bmatrix} \frac{1}{\sigma_1^2 \sigma_2^2} \begin{bmatrix} \sigma_2^2 & 0 \\ 0 & \sigma_1^2 \end{bmatrix} \begin{bmatrix} x_1 - \mu_1 \\ x_2 - \mu_2 \end{bmatrix} \\ &= -\frac{1}{2} \begin{bmatrix} x_1 - \mu_1 & x_2 - \mu_2 \\ \hline \end{bmatrix} \frac{1}{\sigma_1^2 \sigma_2^2} \begin{bmatrix} \sigma_2^2 (x_1 - \mu_1) \\ \sigma_1^2 (x_2 - \mu_2) \end{bmatrix} \\ &= -\frac{1}{2 \sigma_1^2 \sigma_2^2} \left[\sigma_2^2 (x_1 - \mu_1)^2 + \sigma_1^2 (x_2 - \mu_2)^2 \right] \end{aligned}$$

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Now, transpose of cofactor of sigma, this is again coming the same thing because rho and column interchanged symmetry 1, 0, 0, sigma 1 square. Then, what is my inverse, inverse is that is the cofactors mean 1 by determinant, you write down first sigma 1 square, and sigma 2 square that is the determinant sigma 2 square, 0, 0, sigma 1 square.

Now, calculate this one my calculation is minus half, our X 1 minus mu 1 X 2 minus mu 2. Then, inverse is coming like this 1 by sigma 1 square sigma 2 square sigma 2 square, 0, 0, sigma 2 square then multiplied by X 1 minus mu 1 and X 2 minus mu 2. Suppose, if I do this portion first, what you will get you will get minus half X 1 minus mu 1 X 2 minus mu 2 1 by sigma 1 square sigma 2 square this is 2 cross 2 this is 2 cross 1 you will be getting 2 cross 1 this into this plus this into this.

So, it is basically sigma 2 square X 1 minus mu 1 then this into this plus this plus 0 then 0 again, sigma 1 square X 2 minus mu 2. Now, let me bring this one this side later on, we will manipulate sigma 1 square, sigma 2 square, so if you multiply this 1 cross 2 and 2 cross 1, you will be getting 1 cross 1. So, this into this plus this into this you see what is happening sigma 2 square X 1 minus mu 1 square because X 1 minus mu 1 X 1 minus mu 1 plus sigma 1 square X 2 minus mu 2 square. So, if you divide this 2 by sigma 1 square sigma 2 square what you will be getting?

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The image shows a whiteboard with handwritten mathematical derivations. At the top, there is an equation:
$$= -\frac{1}{2} \left[\left(\frac{x_1 - \mu_1}{\sigma_1} \right)^2 + \left(\frac{x_2 - \mu_2}{\sigma_2} \right)^2 \right]$$
 Below this, the joint density function for two variables is derived:
$$f(x_1, x_2) = f(x_1) \times f(x_2) = \frac{1}{(2\pi)^2 |\Sigma|^{1/2}} e^{-\frac{1}{2} \left[\left(\frac{x_1 - \mu_1}{\sigma_1} \right)^2 + \left(\frac{x_2 - \mu_2}{\sigma_2} \right)^2 \right]}$$
 This is then simplified to the matrix form:
$$= \frac{1}{(2\pi)^2 |\Sigma|^{1/2}} e^{-\frac{1}{2} [(x-\mu)^T \Sigma^{-1} (x-\mu)]}$$
 Finally, the general multivariate normal density function is given as:
$$f(x_1, x_2, \dots, x_p) = \frac{1}{(2\pi)^{p/2} |\Sigma|^{1/2}} e^{-\frac{1}{2} [(x-\mu)^T \Sigma^{-1} (x-\mu)]}$$
 with the note $-\infty < x_j < \infty, j = 1, 2, \dots, p.$ There are also small logos for 'CET LIT. NIGP' and 'NPTEL' on the whiteboard.

You will be getting like this minus half, then it will be sigma 1 minus mu 1 by sigma 1 square plus X 2 minus mu 2 by X 1 minus mu 1 by sigma 1 square this one. You have seen that we have found out earlier also this is f x 1 cross f x 2, this we will find out like this one you found out you. Just check I showing that earlier when we have multiplied the two what we got here 1 by 2 pi 2 2 pi to the power 2 by 2 sigma 1 square sigma 2 square half then minus 1 by 2 X 1 minus mu 1 by sigma square this one.

Here, what are you getting here same thing you are getting, so what I mean to mean to say all though this is not a derivation this is the other way proof that what we are saying that means I can write for a bivariate case. I can write my bivariate normal distribution like this is the case you can write like this.

If it is true for multivariate case also then what will happen, ultimately multivariate case X 1, X 2 and X p, then it will be 2 pi to the power p by 2 determinant of this then e to the power minus half x minus mu transpose. This is the case and you have to write minus infinite x j infinite j equal to 1 2 p, this is our multivariate normal distribution we say multivariate normal density function defined. Now, what will be the bivariate density normal density function when your matrix is like this.

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$$\Sigma = \begin{bmatrix} \sigma_1^2 & \sigma_{12} \\ \sigma_{12} & \sigma_2^2 \end{bmatrix} \quad \mu = \begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix} \quad X = \begin{bmatrix} X_1 \\ X_2 \end{bmatrix}$$

$$|\Sigma| = \sigma_1^2 \sigma_2^2 - \sigma_{12}^2 = \sigma_1^2 \sigma_2^2 - (\rho_{12} \sigma_1 \sigma_2)^2 = \sigma_1^2 \sigma_2^2 (1 - \rho_{12}^2)$$

$$\Sigma^{-1} = \frac{1}{\sigma_1^2 \sigma_2^2 - \sigma_{12}^2} \begin{bmatrix} \sigma_2^2 & -\sigma_{12} \\ -\sigma_{12} & \sigma_1^2 \end{bmatrix}$$

$$-\frac{1}{2} (X-\mu)^T \Sigma^{-1} (X-\mu) = -\frac{1}{2} [X_1 - \mu_1, X_2 - \mu_2] \frac{1}{\sigma_1^2 \sigma_2^2 - \sigma_{12}^2} \begin{bmatrix} \sigma_2^2 & -\sigma_{12} \\ -\sigma_{12} & \sigma_1^2 \end{bmatrix} \begin{bmatrix} X_1 - \mu_1 \\ X_2 - \mu_2 \end{bmatrix}$$

This covariance matrix is like this sigma 1 square, sigma 1 2, sigma 1 2, sigma 2 square that means there is covariance this the case what will happen can you not find out this determinant of this is our sigma 1 square, sigma 2 square, minus sigma 1 2 square. So, this can be written like this sigma 1 square, sigma 2 square minus rho 1 2 sigma 1 sigma 2 square covariance is correlation times the standard deviations. So, we can write this one sigma 1 square sigma 2 square 1 minus rho 1 2 square and what will be your inverse here now inverse will be 1 by determinant.

So, 1 by determinant, let me keep this one only then sigma 1 square sigma 2 square minus sigma 1 2 square into we know that transpose of the cofactor. So, I will take this, so it will be sigma 2 square then what will be this plus this is minus sigma 1 2 minus sigma 1 square. Then, what is my exponent part half x minus mu transpose X minus mu transpose sigma inverse x minus mu. This is equal to minus half X 1 minus mu 1 X 2 minus mu 2 X 1 minus mu 1 mu 2 the 1 by sigma 1 square sigma 2 square minus sigma 1 2 square into sigma 2 square minus sigma 1 2 minus sigma 1 2 sigma 1 square times X 1 minus mu 1 X 2 minus mu 2 correct.

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$$\begin{aligned}
 &= -\frac{1}{2} \left[\frac{x_1 - \mu_1 \quad x_1 - \mu_2}{1 \times 2} \right] \frac{1}{\sigma_1^2 \sigma_2^2 - \sigma_{12}^2} \left[\begin{array}{c} \sigma_2^2 (x_1 - \mu_1) - \sigma_{12} (x_2 - \mu_2) \\ -\sigma_{12} (x_1 - \mu_1) + \sigma_1^2 (x_2 - \mu_2) \end{array} \right] \\
 &= -\frac{1}{2} \cdot \frac{1}{\sigma_1^2 \sigma_2^2 - \sigma_{12}^2} \left[\begin{array}{c} \sigma_2^2 (x_1 - \mu_1)^2 - \sigma_{12} (x_1 - \mu_1) (x_2 - \mu_2) - \sigma_{12} (x_1 - \mu_1) (x_2 - \mu_2) \\ + \sigma_1^2 (x_2 - \mu_2)^2 \end{array} \right] \\
 &= -\frac{1}{2} \frac{1}{\sigma_1^2 \sigma_2^2 - \sigma_{12}^2} \left[\begin{array}{c} \sigma_2^2 (x_1 - \mu_1)^2 - 2\sigma_{12} (x_1 - \mu_1) (x_2 - \mu_2) + \sigma_1^2 (x_2 - \mu_2)^2 \end{array} \right] \\
 &= -\frac{1}{2} \frac{\sigma_2^2}{\sigma_1^2 \sigma_2^2 - \sigma_{12}^2} \left[\begin{array}{c} \left(\frac{x_1 - \mu_1}{\sigma_1} \right)^2 - 2 \frac{\sigma_{12}}{\sigma_1 \sigma_2} \left(\frac{x_1 - \mu_1}{\sigma_1} \right) \left(\frac{x_2 - \mu_2}{\sigma_2} \right) + \left(\frac{x_2 - \mu_2}{\sigma_2} \right)^2 \end{array} \right] \\
 &= -\frac{1}{2} \frac{\sigma_2^2}{\sigma_1^2 \sigma_2^2 - \sigma_{12}^2} \left[\begin{array}{c} \left(\frac{x_1 - \mu_1}{\sigma_1} \right)^2 - 2 \rho_{12} \left(\frac{x_1 - \mu_1}{\sigma_1} \right) \left(\frac{x_2 - \mu_2}{\sigma_2} \right) + \left(\frac{x_2 - \mu_2}{\sigma_2} \right)^2 \end{array} \right]
 \end{aligned}$$

So, if you further manipulate, what will happen this equal to minus half X 1 minus mu 1 X 2 minus mu 2, then I want to multiply the last two parts. So, I am writing like this sigma 1 square sigma 2 square minus sigma 1 2 square into this one, you see this is 2 cross 2 and this one is 2 cross 1. So, this multiplied by this plus this multiplied by this plus this multiplied by this, so if we write like this what will happen here sigma 2 square into X 1 minus mu 1 minus sigma 1 2 into X 2 minus mu 2.

So, that is coming from this part row column second one will be minus sigma 1 2 minus sigma 1 2, we are multiplying this with this. So, X 1 minus mu 1 an then plus sigma 1 square when you are saying sigma 1 X 2 minus mu 2, so that is your matrix your this is the first row, this is the second row. So, it is 2 cross 1, so then minus half into sigma 1 square sigma 2 square minus sigma 1 2 square you keep here. Now, you are multiplying this into this so multiply this is 1 cross 2, 2 cross 1 you will get 1, so this into this plus this into this.

So, what you are getting then you are basically getting sigma 2 square X 1 minus mu 1 and X 1 minus mu 1 that is square minus what you are getting this into this so sigma 1 2 into X 1 minus mu 1 and X 2 minus mu 2. So, this into this over second column verse here second row, so it will be minus sigma 1 2 X 1 minus mu 1 X 2 minus mu 2 plus sigma 1 square X 2 minus mu 2 square that is the total. So, if I further manipulate this what I can write sigma 1 square sigma 2 square minus sigma 1 2 square then this is

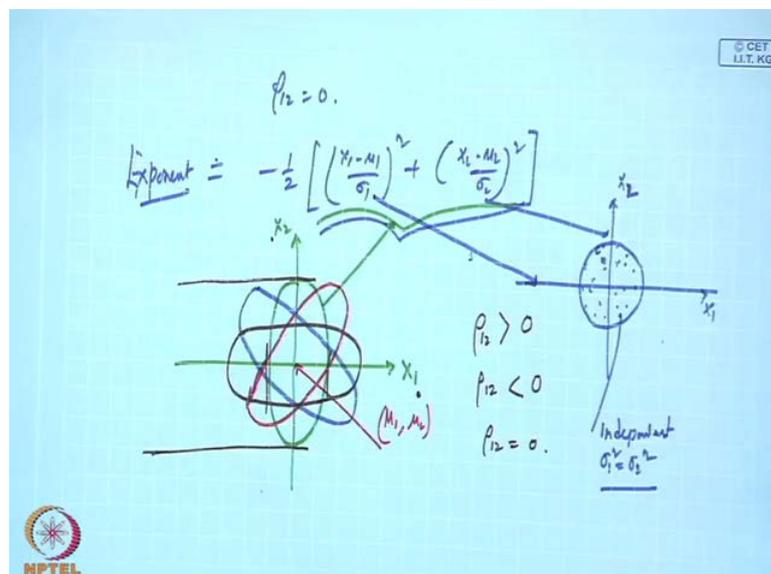
$\sigma_2^2 X_1 - \mu_1^2 - 2\sigma_1\sigma_2\rho(X_1 - \mu_1)(X_2 - \mu_2) + \sigma_1^2 X_2 - \mu_2^2$.

If you divide this within bracket quantity by σ_1^2 and σ_2^2 what will happen $\frac{1}{\sigma_1^2} \frac{1}{\sigma_2^2} \sigma_1^2 \sigma_2^2 \rho$. So, I am dividing the entire thing by σ_1^2 and σ_2^2 I am taking common then what will happen this one $X_1 - \mu_1$ by σ_1 see σ_2 is already there σ_1^2 we have already taken σ_1 I am keeping. Here, this $\frac{1}{\sigma_1^2} \frac{1}{\sigma_2^2} \sigma_1^2 \sigma_2^2 \rho$ then divided by you write σ_1 and σ_2 here can we not write like this, like this, this equal to this X_2 by this see $\sigma_1^2 \sigma_2^2$.

You have taken here plus you can write down $X_2 - \mu_2$ by σ_2^2 what is what is this quantity $\frac{1}{\sigma_1^2} \frac{1}{\sigma_2^2} \sigma_1^2 \sigma_2^2 \rho$, so I can write like this $\frac{1}{\sigma_1^2} \frac{1}{\sigma_2^2} \sigma_1^2 \sigma_2^2 \rho$. You have already seen $\sigma_1^2 \sigma_2^2 \rho^2$ equal to $\sigma_1^2 \sigma_2^2 \rho^2$. So, you take common here $1 - \rho^2$ then this quantity is $X_1 - \mu_1$ by σ_1^2 minus 2ρ $X_1 - \mu_1$ by σ_1 $X_2 - \mu_2$ by σ_2 plus $X_2 - \mu_2$ by σ_2^2 so this quantity this will be cancelled out.

So, if I see this versus the independent case you will very easily find out, now if I put $\rho = 0$, this one will become 0, so then this is $X_1 - \mu_1$ by σ_1^2 plus $X_2 - \mu_2$ by σ_2^2 and what you have here also $\rho = 0$.

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So, I remove $\frac{1}{2}$ by minus half that mean the resultant quantity will be, if I put ρ_{12} equal to 0 my quantity is coming this $\frac{1}{2} X_1^2 - \mu_1$ by σ_1^2 square plus $X_2^2 - \mu_2$ by σ_2^2 square. So, this is the exponent part clear, so that means what I mean to say that in the reverse way also we proved that yes this quantity is following the distribution equal distribution what we have thought of. Now, question comes what is this is the shape of this ellipse correct, now see this diagram this very important concept.

Here, see this is my equation and we have started with this we said this is the scattered plot of X_1 and X_2 and it resembles that there is no dependency between the two variables that mean covariance is 0. We assume σ_1^2 is equal to 0, so that mean this one is nothing but this ellipse what is coming here this ellipse. So, that means if I just write down this one what you are getting you are getting you see that what I will do now I will draw a line like this, but it will be a curve so it is basically coming like this.

So, when you plot this X_1 and X_2 that exponent part you are getting an ellipse when any time you can get an ellipse because this one is also a equation of ellipse please keep in mind this is the in two dimensional general equation of ellipse. So, if I want to plot this what will happen to my figure you are now depending on ρ_{12} yes origin is at μ_1 μ_2 . You usually figure that is original side μ_1 μ_2 data is given in such a manner that 0 is the 0, 0 is the origin μ_1 μ_2 origin is μ_1 and μ_2 this is your μ_1 and μ_2 what will happen if you take the general equation means this 1.

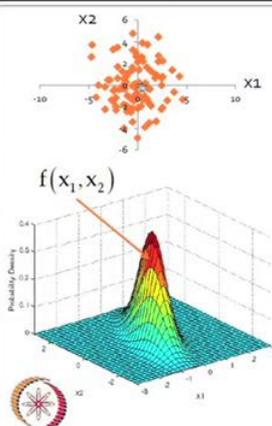
So, depending on the ρ value that ρ_{12} value is it positive is it negative is it 0. If it is 0 this is the diagram this side or you it may because this side see in here the major axis of the ellipse lies along X_2 axis the reason is the variability along X_2 is more they are independent. That is why the major and the minor axis of the ellipse go along the original X_1 and X_2 axis and along X_2 axis the major axis lies because the variability along X_2 axis is more and variability along X_1 is less, sorry variable X_2 is more variable X_1 is less if variability along X_1 is more. Then, they are independent, then your ellipse will become like this keep in mind they are independent when ρ_{12} is greater than 0, it will be so X_1 increases X_2 increases like this so it will be like.

This is inclined because the major and minor axis of the ellipse is not parallel to the original X_1 and X_2 axis. So, as this one is increasing this is also increasing other way

when it is less than it will be just this it will go to this level. In one of the slide I think I have shown you this picture that my X_1 and X_2 is like this my data is like this it is a circle this is also a bivariate case. So, this is also independent case the question is this sigma 1 and sigma 2 in this case independent, but sigma 1 square equal to sigma 2 squares, how do you know this axis. I know the ellipse what is this value suppose this is the first of the entire direction second one is the value how you know all this things.

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Bivariate normal density function



Let x_1 and x_2 are independent with pdf $N(\mu_1, \sigma_1^2)$ and $N(\mu_2, \sigma_2^2)$, respectively. That means $\sigma_{12} = 0$.

So,

$$f(x_1) = \frac{1}{\sqrt{2\pi\sigma_1^2}} e^{-\frac{1}{2}\left(\frac{x_1 - \mu_1}{\sigma_1}\right)^2}; -\infty < x_1 < +\infty \quad \dots(1)$$

$$f(x_2) = \frac{1}{\sqrt{2\pi\sigma_2^2}} e^{-\frac{1}{2}\left(\frac{x_2 - \mu_2}{\sigma_2}\right)^2}; -\infty < x_2 < \infty \quad \dots(2)$$

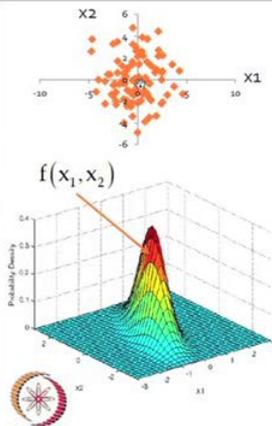
As x_1 and x_2 are independent

$$f(x_1, x_2) = f(x_1) \times f(x_2)$$

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Bivariate normal density function



$$f(x_1, x_2) = f(x_1) \times f(x_2)$$

$$f(x_1, x_2) = \frac{1}{\sqrt{2\pi\sigma_1^2}} e^{-\frac{1}{2}\left(\frac{x_1 - \mu_1}{\sigma_1}\right)^2} \times \frac{1}{\sqrt{2\pi\sigma_2^2}} e^{-\frac{1}{2}\left(\frac{x_2 - \mu_2}{\sigma_2}\right)^2}$$

$$= \frac{1}{(2\pi)(\sigma_1^2\sigma_2^2)^{1/2}} \cdot e^{-\frac{1}{2}\left[\left(\frac{x_1 - \mu_1}{\sigma_1}\right)^2 + \left(\frac{x_2 - \mu_2}{\sigma_2}\right)^2\right]} \quad \dots(3)$$

where, $-\infty < x_1 < \infty$ and $-\infty < x_2 < \infty$.

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Let us see some of the slides here.

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Example-1

■ A process is $X \sim N_2 = (\mu, \Sigma)$ is designed to produce laminar aluminium sheet of length x_1 and breadth x_2 with the following popular parameters (right). Obtain its bivariate normal distribution.

$$\mu = \begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix} = \begin{bmatrix} 100 \\ 50 \end{bmatrix} \text{ and}$$

$$\Sigma = \begin{bmatrix} \sigma_1^2 & \sigma_{12} \\ \sigma_{12} & \sigma_2^2 \end{bmatrix} = \begin{bmatrix} 10 & 0 \\ 0 & 5 \end{bmatrix}$$

$$\frac{1}{44.43} e^{-\frac{1}{2} \left[\left(\frac{x_1 - 100}{\sqrt{10}} \right)^2 + \left(\frac{x_2 - 50}{\sqrt{5}} \right)^2 \right]}$$

Answer :

$$\frac{1}{44.43} e^{-\frac{1}{2} \left[\left(\frac{x_1 - 100}{\sqrt{10}} \right)^2 + \left(\frac{x_2 - 50}{\sqrt{5}} \right)^2 \right]}$$

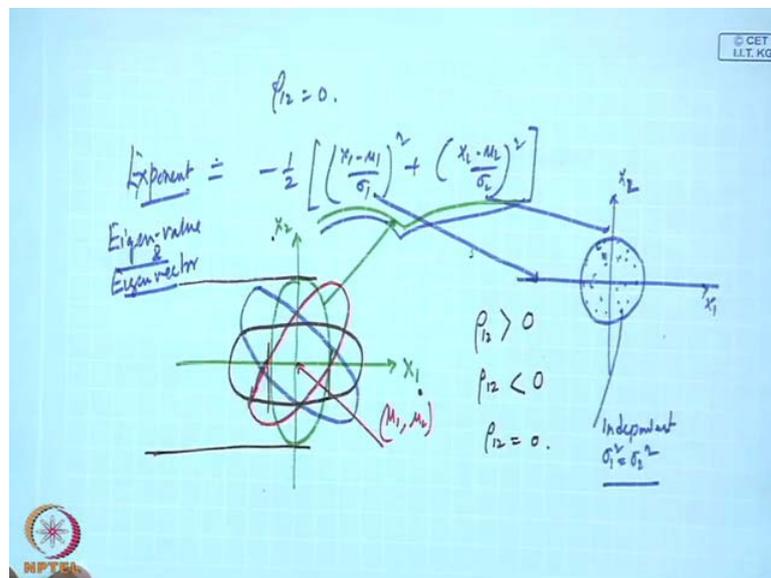


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This is multivariate, so let us see this one first I will I will come back to this how to determine the axis and length all those things. So, what I request to all of you in order to understand the axis.

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You have to know little bit of matrix what is this again value Eigen value Eigen vector, so Eigen value and Eigen vector. I will show you next class Eigen value Eigen vector then axis the all those things and now see one example a process is characterized by two

variables that is X process is designed to produce laminar aluminum sheet of length X 1 and breadth X 2.

With the following population parameters this and this are the population parameters obtain its bivariate normal distribution this is the answer, I am sure you will be able to find out this one from the beginning. If you start the way we have described if you start in the same manner you will ultimately ending. With this answer, we come to properties that multivariate normal distribution has started very, very useful properties, multivariate normal distribution that we will denote.

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$$MND \sim N_p(\mu, \Sigma)$$

$$\mu = \begin{bmatrix} \mu_1 \\ \mu_2 \\ \vdots \\ \mu_p \end{bmatrix} \quad \Sigma = \begin{bmatrix} \sigma_1^2 & \rho_{12} & \dots & \rho_{1p} \\ 0 & \sigma_2^2 & \dots & \rho_{2p} \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \sigma_p^2 \end{bmatrix}$$

$$X_j \sim N(\mu_j, \sigma_j^2), \quad j=1, 2, \dots, p.$$

Next, hence proved that is MND multivariate normal distribution which is we say N p mu and sigma it has many useful properties some of the useful properties I am describing. Now, you see the first property if X is multivariate normal then all the variables individually are invariate normal obvious that when x is there are X 1 to X p. They simultaneously multivariate normal then X 1 is also invariate normal X 2 is invariate normal x p is also invariate normal that means what I mean to say that this one where mu is mu 1 mu 2 mu p and sigma.

I am writing sigma 1 square sigma 2 square sigma p square and these components are also there any one. If say x j this will be your invariate normal with mu j and sigma j square j equal to 1 2 p, so that mean that sigma j will be coming from here that sigma j square.

because that first mu variable you have taken what will be the sigma q sigma. Now, sigma 1 1, sigma 1 2 like sigma 1 q, sigma 1 p, similarly this will be sigma 1 q then somewhere sigma 1 q then sigma q p then sigma 1 p that sigma p p.

So, you have created a subset with q variables, so that means what is happening here, now this is your sigma q and mu case is mu 1 mu 2 mu q mu p, so this is your mu q. So, that means if you take a subset and you know the parameters for those the subset of parameters you consider and find out its distribution that will be multivariate normal distribution the third distribution is very, very useful. The third property is very, very useful, you see what is written if X is if x is multivariate normal linear combination of x j is invariate normal this property can be exploited like anything in your research what is what does it mean?

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$$a^T = [a_1 \ a_2 \ \dots \ a_p]_{1 \times p}$$

$$X = \begin{bmatrix} x_1 \\ x_2 \\ \vdots \\ x_p \end{bmatrix}_{p \times 1}$$

Linear Combination = $a^T X \sim N(a^T \mu, a^T \Sigma a)$.

$$= a_1 x_1 + a_2 x_2 + \dots + a_p x_p.$$

$$E(a^T X) = a^T E(X) = a^T \mu = a_1 \mu_1 + a_2 \mu_2 + \dots + a_p \mu_p.$$

$$V(a^T X) = \frac{a^T \Sigma a}{\frac{1 \times p \quad p \times p \quad p \times 1}{1 \times 1}}$$

MPTEL

It means that suppose I will first create a vector like this a 1 all constant a p some I is my X is X 1 X 2 x p then this one is 1 cross p this is p cross 1 so what is the linear combination, linear combination. Obviously, this one gives you a 1 X 1 plus a 2 X 2 like a p x p what it is say is that the this property says that what will be the expected value of a T x it will basically a T. Expected value of x will be a T and mu this is nothing but a 1 mu 1 a 2 mu 2 a p mu p and what will be your variance of this a transpose x.

This will be you are a transpose sigma a, you can prove it also writing like this, so a transpose sigma. You see it is 1 cross p p cross p p cross 1 resultant is 1 cross 1. So, then

the linear combination will follow univariate normal with a transpose mu a transpose sigma a that is our variance spot. Now, the fourth property fourth property says that instead of one linear combination if you make two linear combinations, what is happening here, you just see in one linear combination.

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Handwritten mathematical derivation on a blue background:

- Top left: $a = \begin{bmatrix} a_1 \\ a_2 \\ \vdots \\ a_p \end{bmatrix}_{p \times 1}$
- Top right: $a^T x = (M)$
- Middle left: $A = \begin{bmatrix} a_{11} & a_{12} & \dots & a_{1p} \\ a_{21} & a_{22} & \dots & a_{2p} \\ \vdots & \vdots & \ddots & \vdots \\ a_{p1} & a_{p2} & \dots & a_{pp} \end{bmatrix}_{p \times p}$
- Middle right: $A^T x \sim N_{a^T \mu, A^T \Sigma A}$ with dimensions $(q \times 1)$, $(q \times p)$, and $(p \times p)$ indicated.
- Bottom: $A^T x = \begin{bmatrix} a_{11}x_1 + a_{12}x_2 + \dots + a_{1p}x_p \\ a_{21}x_1 + a_{22}x_2 + \dots + a_{2p}x_p \\ \vdots \\ a_{p1}x_1 + a_{p2}x_2 + \dots + a_{pp}x_p \end{bmatrix}$

We have taken a 1 2 we have taken a p this is the instead of this I am creating another one like this a 1 1, a 1 2, a 1 p, a 2 1, a 2 2, a 2 p. Suppose, a q 1, a q 2 dot dot dot a q p what is happening now? If I find out, so this one is my 1 2 p, so this is p cross q this one is p cross 1. Now, if I make like this A transpose x what will happen then this will be your q cross p and this will be your p cross 1. So, you will be getting something called q cross 1 where as in one linear combination a cross t is basically 1 cross, so that means q cross 1. This means you are ultimately creating this one q cross 1 a 1 1 X 1 a 1 2 X 2 like this a 1 p x p for this a 2 1 X 1 plus a 2 2 X 2 plus a 2 p x p.

So, like this a q 1 X 1 a q 2 X 2 plus q p x p so if q 1 will be this if I take one combination that is invariate normal take this one second, so all collectively what you are saying collectively they will be multivariate normal. So, that means this quantity will be as q linear combination you have made this into definitely what will happen a sorry A transpose mu. Then, A just check this transpose part you have to check what is A here a is p cross q and this one is q, so that means what do you want this will be cross q if I

write like this. I think in books may they have written in the other way round that part you check ultimate aim is as it is q variable that a transpose x is q variable vector.

So, the variance component will be order of q cross q and mean component be order of q cross 1 column vector definitely. So, this four properties are important and you will you have you see that we you calculate \bar{x} in invariate case. When you calculate \bar{x} that is what that is linear combination of multivariate observations n observations are there 1 by n into x or so equal. Now, then that then what will be the distribution of \bar{x} , although it is invariate normal that is why the sigma square by n is coming there, so all those things. Here, we will be seeing not that \bar{x} only it will be a big \bar{x} that means mean vector for, so next class I will explain you that statistical distance.

Thank you very much.